# R tools to link Game Theory and Statistics by sampling 

Alejandro Saavedra-Nieves

## Motivation: the International Monetary Found

The IMF (January, 2002)


- Analysis of the power of the 179 members when $q=50 \%$.
- A group passes a law when its aggregate weight is larger than $q$.

| Country | Voting rights | Banzhaf-Owen value |
| :---: | :---: | :---: |
| United States | 17.11 | 0.6471 |
| Japan | 6.14 | 0.1752 |
| Germany | 6.00 | 0.1712 |
| France | 4.95 | 0.1411 |
| United Kingdom | 4.95 | 0.1411 |
| Austria | 0.87 | 0.0248 |
| Belarus | 0.19 | 0.0054 |
| Belgium | 2.13 | 0.0608 |
| Czech Republic | 0.39 | 0.0111 |
| Hungary | 0.49 | 0.0140 |
| Kazakhstan | 0.18 | 0.0051 |
| Luxembourg | 0.14 | 0.0040 |
| Slovak Republic | 0.18 | 0.0051 |
| Slovenia | 0.12 | 0.0034 |
| Turkey | 0.46 | 0.0131 |
| Armenia | 0.05 | 0.0014 |
| Bosnia and Herzegovina | 0.09 | 0.0026 |
| Bulgaria | 0.31 | 0.0088 |
| $\ldots$ | $\ldots$ | $\ldots$ |

Alonso-Meijide, J. M., Bowles, C. (2005). Generating functions for coalitional power indices: An application to the IMF. Annals of Operations Research, 137(1), 21-44.

## TU-games

## Game theory: mathematical theory of interactive decision problems

A TU-game is a pair $(N, v)$ :

- $N$ is the set of players and,
$-v: 2^{N} \longrightarrow \mathbb{R}$ is the characteristic function with $v(\emptyset)=0$.


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- $v(S)=1$ (winning), if $S \supset T, T \in\{\{1,2\},\{1,3\},\{2,3\}\}$.
- Otherwise, $v(S)=0$ (not winning).
$S$ is a swing if

$$
v(S \cup i)-v(S)=1, \text { for all } S \subseteq N \backslash\{i\}
$$

That is, $i$ converts $S$ into a winning coalition.

Majority: 34

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A main goal: definition and analysis of rules to allocate $v(N)$ Shapley value

Owen value

Banzhaf value

## Banzhaf-Owen value

Marginal contribution of $i$ to $S: v(S \cup i)-v(S)$, for all $S \subseteq N \backslash\{i\}$.

## The Banzhaf value and the Banzhaf-Owen value

The Banzhaf value for $(N, v)$ is $B z_{i}(N, v)=\frac{1}{2^{n-1}} \sum_{S \subseteq N \backslash\{i\}}(v(S \cup\{i\})-v(S))$, for all $i \in N$.
Let $P=\left\{P_{1}, \ldots, P_{m}\right\}$ be a partition of $N$.
The Banzhaf-Owen value (Owen, 1982) for $(N, v, P)$, for every $i \in N$, is

$$
B z O_{i}(N, v, P)=\sum_{R \subseteq P \backslash P_{(i)}} \frac{1}{2^{m-1}} \sum_{S \subseteq P_{(i)} \backslash\{i\}} \frac{1}{2^{p_{i}-1}}\left(v\left({\underset{P}{l} \in}_{\cup}^{\cup} P_{l} \cup S \cup\{i\}\right)-v\left({\underset{P}{l} \in}_{\cup} P_{l} \cup S\right)\right),
$$

with $i \in P_{(i)} \in P$ and $p_{i}=\left|P_{(i)}\right|$.

- For $i \in N, T \subseteq N \backslash\{i\}$ is compatible with $P$ :

$$
T=\cup_{P_{l} \in R} P_{l} \cup S \text { for } R \subseteq P \backslash P_{(i)} \text { and } S \subseteq P_{(i)} \backslash\{i\}
$$

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Banzhaf, J.F. (1964). Weighted voting doesn't work: a mathematical analysis. Rutgers Law Review 19, 317-343.
國 Owen, G. (1982). Modification of the Banzhaf-Coleman index for games with a priori unions. In: Power, Voting and Voting Power (M.J. Holler, ed.), Physica-Verlag, 232-238.

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## Motivation: a weighted majority game

- $N=\{1, \ldots, n\}$ : the countries.
- $h_{1}, \ldots, h_{n}$ : the weights of $N$.
- $q>0$ imposes the majority.


A weighted majority game is given, for each $S \subseteq N$, by

$$
v(S)=1, \text { when } \sum_{i \in S} h_{i} \geq q . \text { Otherwise, } v(S)=0
$$

- Members are grouped by constituences. Each constituence is an union.

For Austria ( $m=24$ and $p_{i}=10$ )

$$
2^{24-1} 2^{10-1} \text { available compatible coalitions }
$$

Proposal: sampling techniques?

## Estimating the Banzhaf-Owen value

## Problem: estimating a mean on a finite population

Let $(N, v, P)$ be a game with $P=\left\{P_{1}, \ldots, P_{m}\right\}$.

- $B z O_{i}(N, v, P)=B z_{i}\left(N^{*}, v\right)$, with $N^{*}=\left\{k: P_{k} \in P \backslash P_{(i)}\right\} \cup\left\{j: j \in P_{(i)}\right\}$.

A procedure based on sampling techniques
(1) We generate at random a sample $\mathcal{T}=\left\{T_{1}, \ldots, T_{\ell}\right\}$ of $\ell$ coalitions in $N^{*} \backslash\{i\}$.
(2) For each $T_{j} \in \mathcal{T}$,

$$
x\left(R_{j}, S_{j}\right)_{i}=v\left(\cup_{P_{l} \in R_{j}}^{\cup} P_{l} \cup S_{j} \cup\{i\}\right)-v\left(\mathcal{P}_{l} \in R_{j} P_{l} \cup S_{j}\right),
$$

being $R_{j} \subseteq P \backslash P_{(i)}$ and $S_{j} \subseteq P_{(i)} \backslash\{i\}$ such that $T_{j}=\left\{k: P_{k} \in R_{j}\right\} \cup S_{j}$.
(8) The estimation of $\mathrm{BzO}_{i}$ is $\overline{\mathrm{BzO}}_{i}=\frac{1}{\ell} \sum_{j=1}^{\ell} x\left(R_{j}, S_{j}\right)_{i}$, for all $i \in N$.

## Which is the best option of sampling?

## A pseudocode for obtaining coalitions

Take $N$ and an integer $k$ such that $1 \leq k \leq 2^{n}$.
A pseudocode to determine coalitions
(1) Initialize $x=\left(x_{1}, x_{2}, \ldots, x_{n}\right)$.
(2) Take $a=k-1$ and cont $=1$.
(3) Do

$$
x_{\text {cont }}=a-2\left\lfloor\frac{a}{2}\right\rfloor
$$

and update $a=\left\lfloor\frac{a}{2}\right\rfloor$ and cont $=$ cont +1 , being $\lfloor x\rfloor$ the floor function. Repeat Step 3 until cont $>n$.
(4) Finally, the coalition $S$ associated with $k$ is given by

$$
S=\left\{i \in N: x_{i}=1\right\} .
$$

R New function: determine_coalition (k)
sampling coalitions = sampling numbers

```
Sampling with replacement
```

An estimator for the mean under simple random sampling with replacement: $\overline{\mathrm{BzO}}_{i}$

- It is the most usual way of estimating a mean population.
- Theoretical results are well-known for bounding the estimation error.
- We are applying this methodology in a finite population scenario.


Bachrach, Y., Markakis, E., Resnick, E., Procaccia, A. D., Rosenschein, J. S., and Saberi, A. (2010). Approximating power indices: theoretical and empirical analysis. Autonomous Agents and Multi-Agent Systems, 20(2), 105-122.
R- Castro, J., Gómez, D., and Tejada, J. (2009). Polynomial calculation of the Shapley value based on sampling. Computers \& Operations Research, 36(5), 1726-1730.
Sampling
with
replacement

Simple random sampling with replacement


R Proposal 1: sample(1:ntotal, replace=TRUE)
R Proposal 2: library(sampling), with $\operatorname{srswr(ell,ntotal).~}$


An estimator for the mean under simple random sampling without replacement: $\overline{\mathrm{BzO}}{ }_{i}$

- In the exact calculation, each coalition is evaluated only once.
- The hypothesis of non-replacement ensures a lower variance for the estimator (in comparison to the case with replacement).
- It requires the storing of information about the coalitions previously sampled.


Saavedra-Nieves, A., and Fiestras-Janeiro, M. G. (2020). Sampling methods to estimate the Banzhaf-Owen value. Annals of Operations Research (to appear).


Simple random sampling without replacement


P Proposal 1: sample(1:ntotal, replace=FALSE)
R Proposal 2: library(sampling), with srswor(ell, ntotal).


An estimator for the mean under subsampling: $\overline{\overline{B z O}}_{i}$
(1) We take without replacement $\mathcal{R}=\left\{R_{1}, \ldots, R_{\ell_{r}}\right\}$ of $\ell_{r}$ coalitions $R_{j} \subseteq P \backslash P_{(i)}$.
(2) For $R_{j} \in \mathcal{R}$, we take without replacement $\mathcal{S}_{R_{j}}=\left\{S_{j 1}, \ldots, S_{j \ell_{s}}\right\}$ of $\ell_{s}$ coalitions $S_{j k} \subseteq P_{(i)} \backslash\{i\}$.
(3) For $\left(R_{j}, S_{j k}\right)$, we obtain $x\left(R_{j}, S_{j k}\right)_{i}$.
(4) The estimation of $B z O_{i}$ is $\overline{\overline{B z O}}_{i}=\frac{1}{\ell_{r}} \sum_{j=1}^{\ell_{r}}\left(\frac{1}{\ell_{s}} \sum_{k=1}^{\ell_{s}} x\left(R_{j}, S_{j k}\right)_{i}\right)$, for all $i \in N$.

> Sampling without replacement

## Subsampling techniques



R Stage 1: sample(1:ntotal_R, replace=FALSE)
Stage 2: sample (1:ntotal_S, replace=FALSE), for each sampling unit obtained at Stage 1.

- Properties and results on determining sampling sizes are provided in Cochran (2007).


An estimator for the mean under systematic sampling: $\overline{B z O}_{i}^{\text {sys }}$

- The non-replacement hypothesis increases the computational complexity in the approximation.
- We have to store information about those compatible coalitions previously sampled.
- The difficulties in obtaining the variance complicate the task of establishing bounds of the error.

Systematic sampling


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Saavedra-Nieves, A. (2020). A systematic sampling procedure for estimating the Banzhaf-Owen value. Operations Research Letters, 48 (6), 725-731.

## Empirical analysis

- We approximate the Banzhaf-Owen value in examples where it is exactly obtained by specific formulas.
R We compare the Banzhaf-Owen value and the estimations under these approaches.


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## Simulation results

- Similar results in the approximation of the Banzhaf-Owen value with all methodologies.


Figure: Absolute error for the 179 players in the IMF under subsampling (black) and under SRSwor (dashed line).

## Empirical analysis

- Theoretical bounds of the absolute error result conservative in practice.
- Subsampling techniques require less computation time to obtain these estimations.

1000 estimations of the Banzhaf-Owen value: Austria in the IMF

- $\ell=10^{7}$

| Absolute error |  | Minimum | Average | Maximum |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { oे } \\ & \text { in } \\ & \\| \\ & 0 \end{aligned}$ | SRSwr | $3.70 \cdot 10^{-9}$ | $3.90 \cdot 10^{-5}$ | $1.47 \cdot 10^{-4}$ |
|  | SRSwor | $3.70 \cdot 10^{-9}$ | $3.91 \cdot 10^{-5}$ | $1.45 \cdot 10^{-4}$ |
|  | Subs. | $3.70 \cdot 10^{-9}$ | $1.90 \cdot 10^{-4}$ | $7.78 \cdot 10^{-4}$ |
|  | Sys | $1.04 \cdot 10^{-7}$ | $4.27 \cdot 10^{-5}$ | $1.05 \cdot 10^{-4}$ |
|  | SRSwr | $2.80 \cdot 10^{-8}$ | $2.30 \cdot 10^{-5}$ | $1.23 \cdot 10^{-4}$ |
|  | SRSwor | $2.80 \cdot 10^{-8}$ | $2.30 \cdot 10^{-5}$ | $1.23 \cdot 10^{-4}$ |
|  | Subs. | $1.72 \cdot 10^{-7}$ | $1.18 \cdot 10^{-4}$ | $4.64 \cdot 10^{-4}$ |
|  | Sys | $1.28 \cdot 10^{-7}$ | $1.79 \cdot 10^{-5}$ | $5.50 \cdot 10^{-5}$ |
| ¢ | SRSwr | $7.00 \cdot 10^{-9}$ | $3.09 \cdot 10^{-6}$ | $1.76 \cdot 10^{-5}$ |
|  | SRSwor | $7.00 \cdot 10^{-9}$ | $3.08 \cdot 10^{-6}$ | $1.45 \cdot 10^{-5}$ |
|  | Subs. | $9.30 \cdot 10^{-8}$ | $1.59 \cdot 10^{-5}$ | $7.96 \cdot 10^{-5}$ |
|  | Sys | $7.00 \cdot 10^{-9}$ | $2.02 \cdot 10^{-6}$ | $6.91 \cdot 10^{-6}$ |


| Time (sec) |  | Minimum | Average | Maximum |
| :---: | :---: | :---: | :---: | :---: |
| $\begin{aligned} & \text { ò } \\ & \text { in } \\ & 11 \\ & 0 \end{aligned}$ | SRSwr | 218.90 | 227.30 | 245.10 |
|  | SRSwor | 206.47 | 229.19 | 249.27 |
|  | Subs. | 15.67 | 16.12 | 17.61 |
|  | Sys | 197.60 | 213.00 | 238.70 |
| $\begin{aligned} & \text { ஃे } \\ & \text { ㅇ } \\ & \text { ó } \end{aligned}$ | SRSwr | 219.50 | 228.20 | 244.50 |
|  | SRSwor | 210.50 | 225.60 | 241.90 |
|  | Subs. | 7.21 | 8.77 | 10.11 |
|  | Sys | 203.80 | 229.60 | 256.50 |
| $\begin{aligned} & \stackrel{\circ}{\circ} \\ & \stackrel{1}{\infty} \\ & \stackrel{1}{\circ} \end{aligned}$ | SRSwr | 218.90 | 226.90 | 244.40 |
|  | SRSwor | 223.70 | 234.20 | 281.80 |
|  | Subs. | 2.34 | 2.43 | 2.69 |
|  | Sys | 197.70 | 213.10 | 237.70 |

## The Galician milk conflict



## The dairy sector in Galicia*

- The dairy sector generates $1.5 \%$ of GDP
- Dairy farms manage $35 \%$ of the farmland.
- Galicia is the leading dairy power in Spain ( $>50 \%$ of farms and $40 \%$ of production)


## A conflict in Galicia

- It is characterized by high production and low competitiveness.
- The average price is the lowest in Spain.
- European Union regulates it by a system of
 quotas.
*Data source: Consellería de Medio Rural da Xunta de Galicia.Saavedra-Nieves, A., and Saavedra-Nieves, P. (2020). On systems of quotas from bankruptcy perspective: the sampling estimation of the random arrival rule. European Journal of Operational Research, 285(2), 655-669.


## The Galician milk conflict

- After the European milk quotas in March 2015...

| Year | Jan. | Feb. | Mar. | April | May | June | July | Aug. | Sept. | Oct. | Nov. | Dec. |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 2013 | 32.65 | 32.66 | 32.76 | 32.84 | 33.06 | 33.03 | 34.44 | 34.86 | 35.59 | 38.57 | 38.93 | 39.17 |
| 2014 | 39.24 | 38.90 | 38.65 | 36.05 | 35.61 | 35.37 | 33.72 | 33.73 | 33.64 | 32.24 | 32.10 | 31.95 |
| 2015 | 30.52 | 30.60 | 30.30 | 28.80 | 28.40 | 27.90 | 27.60 | 27.70 | 28.30 | 28.70 | 28.70 | 28.80 |

Table: Averaged prices of the milk in Galicia (in euros per 100 litres) in the period 2013-2015.

## How to increase the price of milk?



## A low production scenario

- Reducing the milk production in Galicia?
- If we know a maximum production per municipality, how this reduction affect each of the 190 involved?
- We build a new system of quotas for councils.

[^0]
## A low milk production scenario



Some assumptions

- The total milk production for Galicia decreases with respect to March 2015.
- We assume that the individual capabilities for councils are not reduced.
- Each council intends to maintain the maximum possible quota.


## A low milk production scenario



## A new bankruptcy problem

- Set of agents: the 190 most representative councils.
- Estate: the tons of milk in 2014-2015 for Galicia reduces by $\rho \%, \rho \in(0,100]$
- Claims: the capabilities of milk production (individual production of the councils, March 2015).
*Data source: Consellería de Medio Rural da Xunta de Galicia.


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Computing the Shapley value is a difficult task!
$2^{190}$ coalitions to be evaluated

Proposal: a sampling procedure to estimate the Shapley value

## A low milk production scenario: a case study



## What about variability?

R We approximate the Shapley value for the most representative councils in Galicia.

Top 10 of councils with the largest milk production for Galicia

| Council | A Pastoriza | Lalín | Castro de Rei | Santa Comba | Mazaricos | Chantada | Cospeito | Sarria | Silleda | Arzúa |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Maximum | 52499.22 | 49633.35 | 43188.07 | 40002.56 | 36723.09 | 32731.26 | 32668.26 | 32068.42 | 31319.75 | 30539.21 |
| Average | 52472.49 | 49593.97 | 43161.96 | 39973.33 | 36696.91 | 32708.51 | 32645.75 | 32050.23 | 31300.44 | 30519.81 |
| Minimum | 52425.39 | 49561.72 | 43135.33 | 39942.24 | 36668.44 | 32681.45 | 32626.93 | 32030.53 | 31281.85 | 30502.95 |

Table: Summary of 100 estimations of the milk quotas for the councils with $\rho=40 \%$.

## A low milk production scenario: a case study



## What about variability?

R We approximate the Shapley value for the most representative councils in Galicia.

R We obtain 100 estimations and we do some basic statistics.

- Differences in the results seem bearable.

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## Conclusions and further research

- Sampling methods to estimate coalitional values with a priori unions.

Saavedra-Nieves, A., García-Jurado, I., and Fiestras-Janeiro, M. G. (2018). Estimation of the Owen value based on sampling. In E. Gil, E. Gil, J. Gil, and M. Á. Gil (Eds.), The Mathematics of the Uncertain: A Tribute to Pedro Gil (pp. 347-356). Springer.
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- Theoretical analysis from a statistical point of view.

R Empirical analysis on examples where coalitional values are known.
Future work
Extension to other coalitional values or other allocation procedures.
R Creation of a package for estimating coalitional values.


[^0]:    *Data source: Consellería de Medio Rural da Xunta de Galicia.

